Time-Series Analysis on Multiperiodic Conditional Correlation by Sparse Covariance Selection and Its Computational Method

Tokyo Institute of Technology Tokyo Institute of Technology

Michael Lie Suzuki Taiji

1 Introduction

In this research, we consider optimization problems where the objective is a negative Gaussian log-likelihood with two penalty terms. This minimization problem estimates the sparse inverse of its input covariance matrix; sparse covariance selection [Dem72]. The first penalty of ℓ_1 -norm regularization term ensures a sparse solution, i.e., one with few nonzero entries, while the second penalty enhances block partitions in the parameter space. One of the well-known application of this estimation problem is Markowitz's portfolio selection [Mar52], where the inverse of covariance matrix is needed as its input. We propose to apply (a) Alternating Direction Method of Multipliers (ADMM) algorithm [Boy+11] and (b) Quadratic Approximation Method [Hsi+11] for solving the problem of sparse covariance selection.

$\mathbf{2}$ Problem Setup

Given a set of i.i.d. data of $\mathbf{Y}_t = {\mathbf{y}_{1,t}, \cdots, \mathbf{y}_{n,t}}, t \in$ $\{1,\cdots,T\}$ drawn from p-variate Gaussian distributionof $\mathcal{N}(\mathbf{y}; \boldsymbol{\mu}, \boldsymbol{\Sigma})$. Let the empirical covariance matrix between \mathbf{Y}_i and \mathbf{Y}_j is given by

$$\mathbf{S}_{ij} := \frac{1}{n} \sum_{k,l} (\mathbf{y}_{k,i} - \hat{\boldsymbol{\mu}}_i) (\mathbf{y}_{l,j} - \hat{\boldsymbol{\mu}}_j)^\top,$$

where $\hat{\boldsymbol{\mu}}_i = \frac{1}{n} \sum_k \mathbf{y}_{k,i}$. Consider a stationary-time process such that the multiperiodic inverse covariance matrix \mathbf{X} can be expressed as

$$\mathbf{X} = \underbrace{\begin{bmatrix} \mathbf{X}_{11} & \mathbf{X}_{12} & \mathbf{X}_{13} & \cdots & \mathbf{X}_{1,T} \\ \mathbf{X}_{12}^{\top} & \mathbf{X}_{22} & \mathbf{X}_{23} & \cdots & \mathbf{X}_{2,T} \\ \mathbf{X}_{13}^{\top} & \mathbf{X}_{23}^{\top} & \mathbf{X}_{33} & \cdots & \mathbf{X}_{3,T} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \mathbf{X}_{1,T}^{\top} & \mathbf{X}_{2,T}^{\top} & \mathbf{X}_{3,T}^{\top} & \cdots & \mathbf{X}_{T,T} \end{bmatrix}}_{Tp \text{ columns}} \right\}_{\mathbf{X}_{2}}^{\mathbf{Y}}$$

We assume that \mathbf{X} is a stationary time-process, such that $\mathbf{X}_{i,i+h} = \mathbf{X}_{j,j+h}$ for all i, j. Given some regularization penalty $\lambda_1 > 0, \lambda_2 > 0$, the ℓ_1 and ℓ_2 regular-ized Gaussian MLE for the inverse covariance matrix can be estimated by solving the following regularized *log-determinant* program:

$$\operatorname{argmin}_{\mathbf{X} \succ 0} \left\{ -\ln \det \mathbf{X} + \sum_{i,j} \operatorname{trace} \left(\mathbf{S}_{ij}^{\top} \mathbf{X}_{ij} \right) + \lambda_1 \sum_{i,j} \|\mathbf{X}_{ij}\|_1 + \lambda_2 \sum_{i,j} \sum_{k>i,l>j} \|\mathbf{X}_{ij} - \mathbf{X}_{kl}\|_2^2 \right\} =: \operatorname{argmin}_{\mathbf{X}_{ij} \succ 0} f(\mathbf{X})$$
subject to
$$\mathbf{X}_{i,i+h} = \mathbf{X}_{j,j+h}, \forall i, j.$$
(2.1)

3 Proposed Algorithm (a) ADMM algorithm

Firstly, the objective function f comprises of two parts, $f(\mathbf{X}) \equiv g(\mathbf{X}) + h(\mathbf{X})$, where

$$g(\mathbf{X}) = -\ln \det \mathbf{X} + \sum_{i,j} \operatorname{trace} \left(\mathbf{S}_{ij}^{\top} \mathbf{X}_{ij} \right),$$
$$h(\mathbf{X}) = \lambda_1 \sum_{i,j} \|\mathbf{X}_{ij}\|_1 + \lambda_2 \sum_{i,j} \sum_{k>i,l>j} \|\mathbf{X}_{ij} - \mathbf{X}_{kl}\|_F^2.$$

The first component $q(\mathbf{X})$ is twice differentiable and

strictly convex, while the second part $h(\mathbf{X})$ is convex but non-differentiable.

Then, we can write Equation (2.1) into the following optimization problem. minimize

 $g(\mathbf{X}) + h(\mathbf{Z})$ subject to $\tilde{\mathbf{X}} = \tilde{\mathbf{Z}}$

(3.1)

where

$$\begin{aligned} & \mathbf{\tilde{X}} = \mathbf{\tilde{X}} \\ & h(\tilde{\mathbf{Z}}) = \lambda_1 \sum_{i,j} \|\mathbf{Z}_1\|_1 + \lambda_2 \sum_{i,j} \|\mathbf{Z}_2\|_F^2, \\ & \tilde{\mathbf{X}} = \left[(\mathbf{X}')^\top, (\mathbf{D}\mathbf{X}')^\top, (\mathbf{H}\mathbf{X}')^\top \right]^\top, \\ & \tilde{\mathbf{Z}} = \left[(\mathbf{Z}_1)^\top, (\mathbf{Z}_2)^\top, (\mathbf{0})^\top \right]^\top \end{aligned}$$

for appropriate matrices \mathbf{D} and \mathbf{H} that represent set of time difference matrices and set of stationary time difference matrices, respectively. The augmented Lagrangian of Equation (3.1) is

 $L_{\rho}(\tilde{\mathbf{X}}, \tilde{\mathbf{Z}}, \mathbf{Y}) = g(\mathbf{X}) + h(\tilde{\mathbf{Z}}) + (\rho/2) \left\| \tilde{\mathbf{X}} - \tilde{\mathbf{Z}} + \frac{\mathbf{Y}}{\rho} \right\|_{F}^{2}.$ Therefore, the iteration of ADMM is given as:

Ź

$$\begin{split} \mathbf{X}^{\text{-update:}} & \\ & \tilde{\mathbf{X}}^{(k)} := \operatorname*{argmin}_{\tilde{\mathbf{X}}} \left(g(\mathbf{X}) + \frac{\rho}{2} \left\| \tilde{\mathbf{X}} - \tilde{\mathbf{Z}}^{(k)} + \frac{\mathbf{Y}^{(k)}}{\rho} \right\|_{F}^{2} \right) \\ & \\ & \tilde{\mathbf{Z}}^{\text{-update:}} \\ & \\ & \tilde{\mathbf{Z}}^{(k)} := \operatorname*{argmin}_{\tilde{\mathbf{X}}} \left(h(\tilde{\mathbf{Z}}) + \frac{\rho}{2} \left\| \tilde{\mathbf{X}}^{(k+1)} - \tilde{\mathbf{Z}} + \frac{\mathbf{Y}^{(k)}}{\rho} \right\|_{F}^{2} \right), \\ & \\ & \\ & \\ & \\ & \\ & \\ & \mathbf{Y}^{(k+1)} := \mathbf{Y}^{(k)} + \rho \left(\tilde{\mathbf{X}}^{(k+1)} - \tilde{\mathbf{Z}}^{(k+1)} \right). \end{split}$$

(b) Quadratic Approximation

The summary of the algorithm is given as follows: **Input:** Empirical covariance matrices \mathbf{S} , time difference T,

scalar values of λ_1, λ_2 , inner stopping tolerance ϵ

Output: Minimizer X of f(X). 1: for $k = 0, 1, \cdots$ do

- Compute $\mathbf{W}^{(t)} = (\mathbf{X}^{(t)})^{-1}$. 2:
- Form the second order approximation of $f(\mathbf{X}^{(t)} + \boldsymbol{\Delta})$. 3:
- Partition the variables into free and fixed sets based 4: on the gradient.
- Use coordinate descent to find the Newton direction 5: $\mathbf{D}^{(t)}$ over the free variable set.
- Use an Armijo-rule based step-size selection to get α . 6: Update $\mathbf{X}^{(t+1)}$ $= \mathbf{X}^{(t)} + \alpha \mathbf{D}^{(t)}$

4 Conclusion

Both algorithms show consistency on the optimal values and as we can see from the following table that Quadratic Approximation method runs much faster compared to the ADMM algorithm.

Т	n = 20		n = 30	
	(a)	(b)	(a)	(b)
1	179.65	0.75	283	2.11
2	418.25	1.9	739.43	4.22
3	876.94	5.71	1725.67	12.84

Table 1: Runtime (in seconds) for both algorithms.

References

- [Bov+11]S. Boyd, N. Parikh, E. Chu, B. Peleato, and J. Eckstein. "Distributed optimization and statistical learning via the alternat-ing direction method of multipliers." In: Foundations and Trends ing direction method of multipliers." In: in Machine Learning 3.1 (2011), pp. 1-122.
- [Dem72] A. P. Dempster. "Covariance selection." In: *Biometrics* (1972), pp. 157-175.
- [Hsi+11]C.-J. Hsieh, I. S. Dhillon, P. K. Ravikumar, and M. A. Sustik. "Sparse inverse covariance matrix estimation using quadratic approximation." In: Advances in Neural Information Processing Sys-tems. 2011, pp. 2330–2338.
- H. Markowitz. "Portfolio selection." In: The Journal of Finance 7.1 (1952), pp. 77-91. [Mar52]