

Simultaneous confidence bands for contrasts among several non-linear regression curves

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Spurrier (1999, JASA) constructed exact simultaneous confidence bands for linear regression curves. Jamshidian, et al. (2010, CSDA) proposed a simulation-based method when the explanatory variable is restricted to an interval and the design matrices for each group may be different. Naiman's (1986, AS) gives a method for constructing conservative Scheffé-type simultaneous confidence bands single curvilinear regression model over a finite intervals. We consider constructing simultaneous confidence bands for all contrasts of k non-linear regression models by means of the volume-of-tube method (e.g., Takemura and Kuriki (2002, AAP)).

Suppose that observations (x_j, y_{ij}) are available from k groups, and for each group we assume non-linear regression models

$$y_{ij} = \beta_i^T f(x_j) + \epsilon_{ij}, \quad i = 1, \dots, k, \quad j = 1, \dots, n.$$

Here, random errors ϵ_{ij} are assumed to be independently distributed as the normal distribution $N(0, \sigma(x_j)^2)$, where $\sigma(x)$ is a known function, and $f(x)$ is a $p \times 1$ known vector function. The least square estimator $\hat{\beta}_i$ distributed normally with mean of β_i and covariance matrix Σ , here

$$\Sigma = \left(\sum_{j=1}^n \sigma(x_j)^{-2} f(x_j) f(x_j)^T \right)^{-1}$$

is the inverse of the $p \times p$ information matrix.

Let \mathcal{C} be the set of vectors $c = (c_1, \dots, c_k)^T$ such that $\sum_{i=1}^k c_i = 0$. The focus of this paper is to construct $1 - \alpha$ level simultaneous confidence bands for all the contrasts among the k regression models over a given finite interval $\mathcal{X} = (l, u)$ or infinite interval of the covariate for all $x \in \mathcal{X}$ and $c \in \mathcal{C}$. Specifically, we consider simultaneous confidence bands $\sum_{i=1}^k c_i y_i(x) \in \sum_{i=1}^k c_i \hat{y}_i(x) \pm b_\alpha \sqrt{\text{Var}(\sum_{i=1}^k c_i (\hat{y}_i(x) - E(\hat{y}_i(x))))}$, for all $x \in \mathcal{X}$ and $c \in \mathcal{C}$. Here, $\hat{y}_i(x) = \hat{\beta}_i^T f(x)$, $y_i(x) = E(\hat{y}_i(x)) = \beta_i^T f(x)$.

For this purpose, we need to find $b = b_\alpha$ satisfying the equation

$$P \left(\max_{x \in \mathcal{X}, c \in \mathcal{C}} \frac{|\sum_{i=1}^k c_i (\hat{y}_i(x) - E(\hat{y}_i(x)))|}{\sqrt{\text{Var}(\sum_{i=1}^k c_i (\hat{y}_i(x) - E(\hat{y}_i(x)))}} \geq b \right) = P \left(\max_{q \in \Gamma} \langle \xi, q \rangle \geq b \right) = \alpha,$$

where $\xi = \begin{pmatrix} \xi_1 \\ \vdots \\ \xi_k \end{pmatrix}$, $\xi_i = \Sigma^{-\frac{1}{2}}(\hat{\beta}_i - \beta_i)$, $\Gamma = \{c \otimes \psi(x) \mid x \in \mathcal{X}, c \in \mathcal{C}\}$, $\psi(x) = \frac{\Sigma^{\frac{1}{2}} f(x)}{\|\Sigma^{\frac{1}{2}} f(x)\|} \in \mathbb{S}^{p-1}$. Here \otimes

denotes kronecker product. We calculate this probability by means of the volume of tube method.

Tube formula is given by

$$P \left(\max_{q \in \Gamma} \langle \xi, q \rangle \geq b \right) = \frac{\Gamma(\frac{k}{2})}{\sqrt{\pi} \Gamma(\frac{k-1}{2})} \text{Vol}(M) \{P(\chi_k^2 > b^2) - P(\chi_{k-2}^2 > b^2)\} \\ + P(\chi_{k-1}^2 > b^2) + O \left(b^{n-2} e^{-\frac{1}{2}(1+\tan^2 \theta_c)} \right)$$

as $b \rightarrow \infty$, where $M = \{\pm \psi(x) \mid x \in \mathcal{X}\} \subset \mathbb{S}^{p-1}$, $\text{Vol}(M)$ denotes the length of M , and θ_c is the critical radius of Γ .

The critical radius θ_c of Γ is given by

$$\tan^2 \theta_c = \inf_{(x, \theta) \neq (\tilde{x}, \tilde{\theta})} \frac{(1 - \alpha s)^2}{1 - s^2 - \max\{0, \alpha k\}^2},$$

where $s = \psi'(x) \psi(\tilde{x})$, $\alpha = h'(\theta) h(\tilde{\theta})$, $k = \sqrt{m} t$, $t = \psi'_x(x) \psi(\tilde{x})$ and $m = (\psi_x(x))' \psi_x(x)$. We find that large critical radius makes good accuracy.

Illustrative numerical data analyses to compare several growth curves are demonstrated.